

Driver, C., Temple, P. & Urga, G. (2008). Real options - delay vs. pre-emption: Do industrial characteristics matter?. *International Journal of Industrial Organization*, 26(2), pp. 532-545. doi: 10.1016/j.ijindorg.2007.03.003



**CITY UNIVERSITY  
LONDON**

[City Research Online](#)

**Original citation:** Driver, C., Temple, P. & Urga, G. (2008). Real options - delay vs. pre-emption: Do industrial characteristics matter?. *International Journal of Industrial Organization*, 26(2), pp. 532-545. doi: 10.1016/j.ijindorg.2007.03.003

**Permanent City Research Online URL:** <http://openaccess.city.ac.uk/6958/>

### Copyright & reuse

City University London has developed City Research Online so that its users may access the research outputs of City University London's staff. Copyright © and Moral Rights for this paper are retained by the individual author(s) and/ or other copyright holders. All material in City Research Online is checked for eligibility for copyright before being made available in the live archive. URLs from City Research Online may be freely distributed and linked to from other web pages.

### Versions of research

The version in City Research Online may differ from the final published version. Users are advised to check the Permanent City Research Online URL above for the status of the paper.

### Enquiries

If you have any enquiries about any aspect of City Research Online, or if you wish to make contact with the author(s) of this paper, please email the team at [publications@city.ac.uk](mailto:publications@city.ac.uk).

# Real Options - Delay or Pre-Emption: Do Industrial Characteristics Matter?

Ciaran Driver

*Tanaka Business School, Imperial College, London (U.K.)*

Paul Temple

*Department of Economics, University of Surrey (U.K.)*

Giovanni Urga \*

*Cass Business School, London (U.K.)*

March 21, 2007

---

\**Corresponding author:* Faculty of Finance, Centre for Econometric Analysis, Cass Business School, 106 Bunhill Row, London EC1Y 8TZ (U.K.). Tel. +/44/20/70408698; Fax. +/44/20/70408881, e-mail: g.urga@city.ac.uk; www.cass.city.ac.uk/faculty/g.urga. We wish to thank Robert Chirinko, Vivek Ghosal and Hashem Pesaran for perceptive comments. The usual disclaimer applies. J. Munoz Bugarin and K. Imai provided research assistance. ESRC funding under grant R022250159 is gratefully acknowledged.

# **Real Options - Delay or Pre-Emption: Do Industrial Characteristics Matter?**

## **ABSTRACT**

This paper presents an empirical study of the channels of influence from uncertainty to fixed investment suggested by real options theory. Using panel data from the Confederation of British Industry (CBI) Industrial Trends Survey, we report OLS estimates of the impact of uncertainty on investment where the regressors are augmented by cross-sectional averages of the dependent variable and of the individual specific regressors, as recently suggested by Pesaran (2006). The cross-industry pattern of results is checked for consistency with the pattern predicted by real options theory, using a specially constructed data set of industrial characteristics. We find that irreversibility is able to predict the pattern detected, but only when combined with a measure of the information advantage of delay. There is also evidence for expansion options effects; industries with high R&D and advertising intensities tend to have positive uncertainty effects.

***J.E.L. Classification Numbers:*** *E22, C23*

***Keywords:*** *Investment, Industry, Irreversibility, Real Options, Uncertainty.*

# 1 Introduction

Theoretical developments over several decades have highlighted the potential significance of uncertainty for capital investment decisions. A standard argument is that uncertainty should raise the amount of investment because of the likely convexity of marginal profit in the uncertain variable working through Jensen's inequality (Abel 1983). However, traditional convexity models are subject to the critique that they often ignore irreversibility and the timing decision associated with a project. Real options theory provides one explanation for a delayed response under uncertainty to signals that would cause entry or exit in a frictionless world. In this paper we investigate the empirical validity of this approach.

In real options theory, the trigger values for irreversible investment or disinvestment are respectively above and below the corresponding Marshallian values (variable cost plus the servicing of sunk cost of entry or exit) under uncertainty, as long as information arrives stochastically over time and waiting is not too costly. Models of the relationship of adjustment speed to uncertainty for irreversible investment are developed in Brennan and Schwartz (1985), McDonald and Siegel (1986), and Dixit and Pindyk (1994). A similar approach, reconciling the theory with standard q-theory of investment, is developed in Abel and Eberly (1994) where it is shown that the extent of the zone of inaction with respect to the forcing variable depends on the level of uncertainty; furthermore, activity outside the zone of inaction is slowed by heightened uncertainty. While most of the literature is concerned with the option to wait, and its effect on delaying investment, under some circumstances, increased uncertainty can accelerate project development, particularly where there is a time to build or where first mover advantages are significant (Bar-Ilan and Strange 1996; Weeds 2002; Folta and O'Brien 2004). The theoretical rationale for this ambiguity in the real option effect is explained in Abel *et al* (1996) where the effect of irreversibility (no downward adjustment) is allied with lack of expandability (no upward adjustment). Thus, the firm's options

consist of a set of both call options and put options<sup>1</sup>. Much of the empirical literature finds a negative relationship between uncertainty and investment. Indeed, this has become something of a stylized fact. Many of these studies report results for *aggregate* investment. However, in this contribution we find that there is considerable heterogeneity in the response of investment to uncertainty across industries.

The main aim of this paper is to exploit the observed heterogeneity across industries in order to examine the relevance of real options in explaining the pattern. The method we adopt is first to establish estimates of both the sign and the magnitude of the impact of uncertainty on investment for each industry. We then use measures of industry characteristics – as suggested by real options theory – to explain the pattern observed. In Section 2 we discuss real options models of investment and the implications for investment decisions under uncertainty. Section 3 introduces the basic investment model used in the paper and reports the results of estimating the model with indicators of uncertainty. Section 4 assesses whether the cross-sectional pattern of uncertainty coefficients can be reconciled with the theory discussed in Section 3. Section 5 concludes.

## 2 The Theoretical Effect of Uncertainty on Investment Demand

Theory suggests a number of possible channels of influence running from uncertainty to investment. In this section we focus on two opposing influences predicted by real options theory (Trigeorgis 2003).

---

<sup>1</sup>Few studies have examined the combined effects of call and put options. Exceptions include Bulan (2005), who interprets irreversibility as a measure of the difference between the value of call options and put options, and Shaanan (2005), who recognises the issue but abstracts from it in estimation.

## 2.1 Deferment options and convexity

The basic “real options” approach stresses an additional cost to investment which attaches to any early exercise of an option to invest. By deferring the project and keeping the option open, costly mistakes may be avoided (Dixit and Pindyck 1994). The idea is very general: because of proprietary assets in knowledge, competences, or spare land, firms may choose the timing of their investment by balancing any loss from delay against the value of extra information that arrives over time. This may explain how uncertainty might raise the hurdle rate and delay investment projects.<sup>2</sup> Note that the empirical relevance of the argument stems from the existence of *both* irreversibility, and some feature of the firm’s environment which makes delay valuable.

## 2.2 Expansion and compound options

The option to delay cannot always be presumed to exist, e.g. in industries characterized by first mover advantages (FMA). Under some circumstances, investment may be speeded up if other influences are favorable. For example in high-technology industries characterized by patent races, uncertainty may increase the value of the option obtained through early investment. Thus,

---

<sup>2</sup>One criticism of this argument is that although the hurdle rate may be raised, so too may the probability of hitting the hurdle with ambiguous implications for investment. Simulation results in Sarkar (2000) for a single firm partial equilibrium model suggest that a positive effect of uncertainty is possible at low levels of uncertainty. It is not clear however whether this can be generalised to the industry case. A further potential criticism of the real options argument is that with perfectly elastic demand (and constant returns to scale), irreversibility is irrelevant since the marginal rate of return on capital is, in these circumstances, invariant to the quantity of capital installed (Caballero 1991). This model effectively neutralises irreversibility through the focus on individual firm price uncertainty with no linkages from investment in one period to the investment decision in the next. However, at the industry level, new entry can erode excess profit and, in conjunction with irreversibility, create an asymmetry in price that biases investment downward by lowering the expected realised price. Thus, “...industry-wide uncertainty will affect irreversible investment by a competitive firm with constant returns to scale much as it would a non-competitive firm or a firm with decreasing returns to scale” (Pindyck 1993, p.274). Evidence in favour of real option effects may be found in Moel and Tufano (2002).

the literature on real options does not unambiguously predict the sign of the uncertainty-investment relationship. Today, several different kinds of options are routinely identified, leading to contradictory effects on investment (Copeland and Antikarov 2001). First, obtaining an option may in itself be a key part of each investment process, creating “compound options” where obtaining an option on an option is a key element in decision making. An example would be where follow-on products can more easily be launched on the back of a first success. Similarly, expansion options confer the ability to respond to higher than expected demand and are important in cases where lead times or adjustment cost would otherwise imply cost penalties. A variety of models have addressed the question of strategic decision-making in a real option framework and in particular whether the existence of FMA not only destroys the option to defer but actually speeds up investment due to the operation of expansion and compound options (Bar-Ilen and Strange 1996; Mason and Weeds 2001; Boyer *et al* 2004; Smit and Trigeorgis 2004).

The most likely industries where durable FMA exist are those with high product R&D intensity and where switching costs are also high - as they may be when firms advertise intensively. High technology and heavily branded goods will also tend to have high quasi-rents and this will give enhance the value of expansion options.

### **2.3 Testing the importance of the two channels**

As noted above different industrial characteristics predict which real option influence is likely to operate in each case. We first identify which industries are affected (positively or negatively) by uncertainty and then, in a second stage, test whether the industrial characteristics can discriminate accurately as to which, if either, influence is present. The industry characteristics are of course only proxies for theoretical variables that feature in the two real option models. The theoretical variables, the predicted impact of uncertainty, and the industrial characteristics used to proxy the theoretical variables are

detailed in Table I. We postpone to Section 4 the measurement of the proxy variables.

[Insert Table I about here]

### **3 Estimates of the impact of uncertainty on investment**

Our modeling strategy requires a reasonably large cross-section of industries and, given the likely presence of lags of the dependent variable, a long time series. Official data series are not sufficiently disaggregated, at least for a sufficient length of time, for current purposes, but a useful alternative are the data on investment authorizations for over forty industries and eighty quarters, publicly available from the Industrial Trends Survey of the Confederation of British Industries (CBI). This is a quarterly survey based upon the replies from an average of over 1000 enterprises. The results are reported at an industrial level. It proved possible to obtain the necessary data for the analysis of 37 industries. Further details of our use of the Survey can be found in the Appendix, Section A1.

The dependent variable is constructed from Question 3b of the Survey which asks respondents: “Do you expect to authorise more or less capital expenditure in the next twelve months than you authorised in the past twelve months on: plant and machinery?” (The possible choices here are ‘More’, ‘Same’ or ‘Less’). The resulting investment authorisations data record the percentage of firms in each industry responding in the three categories. These data differ of course from conventional gross investment data but may in fact offer some advantages (for example, gestation lags can be dispensed with). In practice however, these two variables are linked by a well determined realization function (Driver and Moreton 1992; European Commission, 1997. See also Lamont 2000). Although the fact that our data are qualitative



represents a potential disadvantage, a well established and useful practical result for qualitative data is that the balance (the percentage responding more minus the percentage responding less) is closely correlated with rates of change (Smith and McAleer 1995; Driver and Urga 2004; Butzen *et al* 2003). Accordingly, we denote the investment authorisation balance as  $Auth$ , to represent investment growth. Summary statistics for this and all the variables described below and used in our investment equation can be found in Table A2 in the Appendix.

Our specification for investment authorisations for each industry  $i$  at time  $t$  ( $Auth_{i,t}$ ) in equation (1) below follows the standard accelerator-type specification (Blanchard and Fisher 1989; Berndt 1990 equation 6.14). The accelerator form is chosen over the more common  $q$ -form in the light of the many findings of "fragile or implausible results" with Euler-equation specifications (Mairesse *et al* 1999, p.6). The basic equation is modified in four main ways as discussed below. First, it includes forward looking expectations derived from the Survey as recommended in Chirinko (1993). Second, additional regressors in the form of uncertainty and a measure of financial constraints are included as is now standard in the literature (Chatelain 2003). Third, the estimating equation has the form of an equilibrium correction model, a feature of many investment equations (Butzen *et al* 2003). Finally, to control for unobserved common factors such as the change in the tax regime imposed in the UK in 1984, the equation is augmented by the unweighted arithmetic cross sectional mean ( $\bar{x}_t = \sum_{i=1}^N x_{it}/N$ ) of all the industries in the sample for each right hand side variable. This was recently advocated by Pesaran (2006). The estimating equation is therefore:

$$\begin{aligned}
Auth_{it} = & b_{i,0} + b_{i,1}Auth_{i,t-1} + b_{i,2}Auth_{i,t-2} + b_{i,3}y_{i,t} + b_{i,4}y_{i,t-1} + b_{i,5}y_{i,t}^e + \\
& b_{i,6}opt_{it} + b_{i,7}unc_t + b_{i,8}unc_{i,t-1} + b_{i,9}unc_{i,t-2} + b_{i,10}fi_{i,t-1} + \\
& + b_{i,11}cu_{i,t-1} + b_{i,12}dcu_{i,t} + cross-sectional\ means + e_{i,t}
\end{aligned} \tag{1}$$

where  $i = 1, \dots, 37$  and  $t = 1978 : 1 - 1990 : 2$ . This equation was estimated by OLS for each of the 37 industries using the common specification above but with the lag structure of the focal uncertainty variables tested down one at a time.

In line with the standard specification, our specification includes both lags on the dependent variable, (*Auth*), and on the past output growth term ( $y$ ). The survey data also allow us to construct an expected future growth variable ( $y^e$ ) which represents short period expectations and thus complements the usual accelerator term. Details concerning the precise construction of both  $y$  and  $y^e$  can be found in the Appendix, Section A1. A further forward looking variable, *opt*, reflecting survey-based business confidence is also included. Previous work has shown that this variable captures demand, interest rates and exchange rate influences (Junankar 1989).

The equation also contains a vector of additional terms reflecting both uncertainty, *unc*, and the possibility that firms may be experiencing financial constraints on investment, *fi* (Chatelain 2003). Our uncertainty variable *unc* is derived from the cross-sectional dispersion of beliefs across firms in an industry about prospects for that industry. Assuming a high degree of homogeneity in demand conditions within the industry, the cross-section dispersion of beliefs about the same sector may be regarded as a measure of uncertainty. The precise measure used is the concentration of responses to the survey question on industry optimism<sup>3</sup>. We compute this measure as the entropy (negative concentration) of the three replies (more/same/less). Writing  $S_j$  for the share of reply  $j$  ( $j = 1, 2, 3$ ) we define:  $unc_{it} = \sum_{j=1}^3 [-S_{jit} \log S_{jit}]$ .

---

<sup>3</sup>Uncertainty in real option models is generally captured by the volatility of some key variable. However, it is not always simple to measure volatility. GARCH models can be used to estimate conditional volatilities but convergence is often a problem and in our case we wished to retain as full a sample of industries as possible. Furthermore it can be argued that it is the future path of conditional volatility that is important (Leahy and Whited, 1996) so that our measure, which is based on forward expectations, is particularly appropriate in this regard. Guiso and Parigi (1999) have used Italian data with similar belief dispersion.

The constructed measure is not highly correlated with the level of optimism: the mean absolute value of the correlation coefficient over our sample of industries is 0.13. An even spread in the replies (each share  $S_j$  equal to one third) corresponds to maximum entropy and maximum uncertainty<sup>4</sup>. The entropy variable has been used successfully in other contexts involving surveys with three possible replies to measure the extent of disagreement among respondents (Fuchs, Krueger and Poterba 1998). Using lack of consensus as a measure of uncertainty receives empirical support in a number of studies (Zarnowitz and Lambros 1987, Bomberger 1999).

Our estimating equation also allows for the potential role of financing constraints  $fi$  by using the responses to question 16(c) of the Industrial Trends Survey which allows for both internal and external constraints as one of the reasons for limiting investment authorisations. After experimentation, our preferred measure sums the proportions of respondents who report either a “shortage of internal finance” or an “inability to raise external finance.”

The estimating equation includes an equilibrium correction term in the form of a lagged capacity utilization variable directly recorded in the Survey ( $cu$ ). Such utilization terms often appear in investment equations to capture (integral control) cumulative deviation from target levels - in this case the capital-output ratio. Given this specification, it is also standard to include the first difference of the  $cu$  ( $dcu$ ) as the dynamic counterpart to the levels term. Further details of the construction of these variables is given in the

---

<sup>4</sup>The *unc* variable may be measured with error. However, using the standard Hausman test procedure, we rejected the hypothesis that OLS estimates were statistically different from IV counterparts.

A further possible criticism of this uncertainty measure is that respondents may mistakenly reply to the survey question by projecting forecasts for their own firm on to the industry as a whole so that the spread of replies on industry optimism becomes an indicator of objective diversity. However the question posed in the survey is quite explicit on this point. Furthermore, we find that the entropy of optimism (relating to the industry) is significantly less than the entropy of output (relating to the firm) in all but four of the industries. We take this as evidence that firms are not just looking at their own fortunes in answering the optimism question.

Appendix.

The final vector of variables in the model consists of the controls for unobserved common factors, i.e, the cross section averages discussed above.

A summary of results from estimating (1) are reported in Table II. The *unc* coefficients are shown in standardised form to assist cross-industry comparisons<sup>5</sup>. Alongside the coefficients (which are shown in separate columns depending on the lag structure identified) we show the level of significance of the tested-down models using a minimum (one-sided) significance level of 10% that allows us to partition the industry set into negative, positive and null uncertainty effects<sup>6</sup>. In the two cases where more than one uncertainty term is retained, an *F* – *test* on the joint significance of the coefficients is reported in column 10. The overall equation diagnostics are summarised in the final three columns and indicate that all equations are well determined. The *F* – *statistic* (column 14) which tests the hypothesis that all the coefficients (except the constant) are zero is significant for all the industries at the 1% level or better.

The results do not support a simple pattern of a negative relationship of investment to uncertainty as suggested by much of the current literature, e.g. that cited in Carruth *et al* (2001). Instead we find a range of coefficients, from positive to negative, though a substantial number of the values are insignificant. The purpose of the next stage of analysis in the following section is to examine the predictability of this pattern.

**[Insert TABLE II about here]**

---

<sup>5</sup>These standardised coefficients are obtained by multiplying the raw coefficient by the standard deviation of *unc* in each industry and then normalising by the standard deviation of the dependent variable.

<sup>6</sup>The full set of results with all variables is available on request. From these it is apparent that for nearly all industries, at least one of the cross-sectional averages of the dependent variable and its lags is significant at the 5% level. Additionally, in the majority of cases, at least one extra average is also significant. For completeness we tested for the joint significance of all the cross-sectional averages and found them to be significant or borderline significant in about a third of cases.

## 4 Option values and the pattern of uncertainty effects on investment

To what extent can real options theory explain the pattern of coefficients contained in Table II? The table indicates 11 industries with significant *unc* effects at 10% or better, of which 7 are significantly negative. These are for ferrous metals, building materials, metal goods, constructional steel-work, electrical industrial goods, electrical consumer goods, and clothing and fur. Four are significant and positive (agricultural machinery, electronic consumer goods, aerospace, and wool textiles).

Our methodology requires us to construct an econometric model of the pattern of results described above. As discussed earlier, if the real options effect is important then the magnitude and sign of the uncertainty effect depends on the balance of the value of the deferment and expansion options as summarized in Table I above. To be able to test for the former, we need to measure the irreversibility associated with investment in any particular industry. For the latter, we require an indicator of the opportunities that will follow on from first-stage investments or indicators of the value of expansion options.

The measure of irreversibility (*irr*) is based upon a ranking of the ratio between second hand plant and equipment sales to the acquisition of such assets, averaged over an economic cycle. Where second-hand markets are thin, this ratio will be close to zero. Further details of all variables used in this analysis can be found in the Appendix, Section A3 and Table A4.

The option to wait will be more valuable when the random process determining investment decisions is highly persistent. Although mean reverting behaviour does not destroy option value it will reduce it (Sarkar 2003). Accordingly we also develop a measure of the persistence of the process, which in our case is calculated from the optimism variable (*opt*) used in the investment equations and discussed above. It is based on the normalised variance

ratio (Cochrane 1988; Proietti 1996).

$$V_k = (1/k)(Var(opt_t - opt_{t-k})/Var(opt_t - opt_{t-1})) \quad (2)$$

where  $k$  is the chosen lag length (20 quarters in our case)<sup>7</sup>. We call this variable *persis\_opt*. However, the theory of real options suggests that *irr* and the measure of persistence should be seen as interactive. Accordingly, we constructed a new variable, based upon the joint distribution of the two variables, which combines *irr* with our measure of persistence into an augmented measure of irreversibility, *irraug*. This used the quartiles of the distribution of the two variables, attaching the highest score to industries which were in the highest quartile on both measures (= 6). Those in the lowest quartile on both variables had a zero score.

Turning now to the measurement of opportunities for expansion, we base our indicator on both the R&D and advertising intensity of the industry. An industry which has a low intensity of either characteristics is assigned a score of zero; a score of one is assigned to industries exhibiting a high intensity for just one characteristic; industries which are both R&D and advertising intensive are assigned a score of two. This indicator variable is denoted *rdad*.<sup>8</sup> As well as representing expansion options, a high score on *rdad* indicates possible preemption where competing technologies and brands are engaged in winner-takes-all competition. The option to wait would not exist in such circumstances<sup>9</sup>.

In order to consider whether real options can predict the pattern of the

---

<sup>7</sup>Given that fixed investment assets are long-dated, we believe that a five year horizon is reasonable. We tested for sensitivity to this horizon by using a lag length of three years, but it made negligible differences to the results.

<sup>8</sup>The R&D and advertising variables are drawn from a study of EU industry by Davies and Lyons (1996, Table A2.1), who divided industries into high and low R&D and advertising intensities on the basis of the ratios of their expenditure on these variables to sales.

<sup>9</sup>We also constructed a profitability index as a rate of profit (*nprtea*) on total capital installed, adjusted for depreciation (See Appendix, Section A3). However, as this was not significant in any of the equations reported below we do not discuss it further.

uncertainty-investment relationships, we require measures of this pattern as the dependent variable. While a variety of such measures are possible, we consider first the simplest case, where the pattern of coefficient significance in Table II is assigned with  $-1, 0, 1$  outcomes depending upon statistical significance and the predicted sign of the effect ( $-1$  indicating a statistically significant and negative coefficient at the 10% level). The result is the variable *OPROB1*. Alternative dependent variables are then considered as a robustness check.

Table III reports some experiments with these variables. The first four results report experiments using *OPROB1* with an ordered probit estimator. By itself, the measure of irreversibility, *irr*, has no explanatory power (column 1). When the value of waiting – measured by *persis\_opt* – is added however, as suggested by our discussion above, both variables are correctly signed and significant at the 10% level (column 2). Moreover, the augmented measure of irreversibility *irraug* – is significant at the 5% level (column 3), and when *rdad* is included this is also correctly signed and significant (column 4).

The importance of *irraug* appears robust to a number of alternative dependent variables and specifications. In column 5, we use a modified version of *OPROB1* where the pattern of coefficient significance is assigned values  $-2, -1, 0, +1, +2$ , where, for example  $-1$  is negative significance at 10% and  $-2$  is negative significance at 5%. The overall level of significance is now below 5% (0.023). This variable is labelled *OPROB2*.

The previous analysis has used conventional levels of statistical significance of the ordered probit regression to derive the dependent variable. An alternative is to use the estimated magnitudes of the uncertainty impact. Accordingly, column 6 shows results when we use the strength of the uncertainty effect as a dependent variable (*OSUM*). This variable takes on values  $-2, -1, 0, 1$ , or  $2$ . Industries with insignificant uncertainty impacts take on a zero value. The other 11 industries were assigned values according to

sign and whether the sum of the standardized coefficients is greater than the average of these sums across the same 11 industries. Using this approach, Column 6 of Table III reports results which show that both *irraug* and *rdad* are significant at the 5% level<sup>10</sup>.

[TABLE III about here]

## 5 Conclusions

We have estimated a set of investment authorisation equations that are well determined and with acceptable diagnostics using OLS augmented by cross-sectional averages to control for unobserved common effects. Our main interest is in the sign, significance and magnitude of the uncertainty coefficients. In a second stage estimation we used this information to construct a set of limited dependent variables that indicate the importance and sign of the uncertainty effects by industry. These limited dependent variables are then regressed on a specially constructed set of industrial characteristics using the ordered probit model.

Our overall conclusion is that the industries showing positive or negative effects from uncertainty to investment are not random draws; in particular two strong conclusions are evident from the second stage regressions. First, in keeping with the predictions from real (deferment) options, irreversibility is a predictor of a negative effect from uncertainty, but only when combined with a measure of the value of waiting. Secondly, there is evidence that indicators of first mover advantages - such as that provided by R&D and advertising intensity - offset the irreversibility effect and contribute to explaining a positive effect of uncertainty on investment in some industries.

---

<sup>10</sup>These results could be replicated when all 37 industries were assigned an uncertainty impact (rather than just the 11 significant instances) using the sum of the point estimates of the coefficients in an unrestricted model i.e. when all uncertainty coefficients were retained.



These results are robust to using different categorisations of the importance of the uncertainty effects.

## REFERENCES

- Abel, A.B., 1983, "Optimal investment under uncertainty", *American Economic Review*, 73, 228-33.
- Abel, A.B., A.K. Dixit, J.C.Eberly, and R.S.Pindyck, 1996, "Options, the value of capital and investment", *Quarterly Journal of Economics*, August, 753-777.
- Abel, A.B., and J. C. Eberly, 1994, "A unified model of investment under uncertainty", *American Economic Review*, 1369-1384.
- Bar-Ilan, A., and W.C.Strange, 1996, "Investment lags", *American Economic Review*, 86, 610-22.
- Berndt, E.R.,1990, *The Practice of Econometrics: Classic and Contemporary*, Massachusetts: Addison Wesley.
- Brennan, M.J., and E.S. Schwartz, 1985, "Evaluating natural resource investments", *Journal of Business*, 58, 35-157.
- Blanchard, O. J. and S. Fischer, 1989, *Lectures on Macroeconomics*. Cambridge Massachusetts: The MIT Press.
- Bomberger, W.A.,1999, "Disagreement and uncertainty –reply", *Journal of Money Credit and Banking*, 31, 273-276.
- Boyer, M., E Gravel, and P Lasserre, 2004, "Real options and strategic competition: a survey", mimeo, Economics, Univ. de Montreal.
- Bulan, L., 2005, "Real options, irreversible investment and firm uncertainty: New evidence from U.S. firms", *Review of Financial Economics*, 14, 255-279.
- P. Butzen, P., Fuss, C. and Vermeulen, P. 2003, "The impact of uncertainty on investment plans: a survey-data approach" with comment by Ulf von Kalkreuth in P. Butzen and C. Fuss (eds) *Firms' Investment and Finance Decisions: Theory and Empirical Methodology*, Edward Elgar.
- Carruth, A., A. Dickerson, and A Henley, 2000, "What do we know about investment under uncertainty?", *Journal of Economic Surveys*, 14, 119-153.

- Caballero, R.J., 1991, "On the sign of the investment-uncertainty relationship", *American Economic Review*, 81, 279-88.
- Chatelain, J.-B., 2003, "Structural modelling of financial constraints on investment: where do we stand?" in P. Butzen and C. Fuss (eds) *Firms' Investment and Finance Decisions: theory and empirical methodology*, Edward Elgar.
- Chirinko, R. S., 1993. "Business Fixed Investment Spending: Modeling Strategies, Empirical Results, and Policy Implications," *Journal of Economic Literature*, 31, 1875-1911.
- Cochrane, J. H., 1988, "How big is the random walk component in GDP?", *Journal of Political Economy*, 96, 893-920.
- Copeland, T., and V. Antikarov, 2001, *Real Options: a Practitioner's Guide*, New York and London, Texere.
- Davies, S., and B.R. Lyons, 1996, *Industrial Organisation in the European Union: Structure, Strategy*, Oxford: Oxford University Press.
- Dixit, A., and R. Pindyck, 1994, *Investment under Uncertainty*, Princeton: Princeton University Press.
- Driver, C., K. Imai, P. Temple, and G. Urga, 2004, "The effects of uncertainty on UK investment authorization: homogeneous vs heterogeneous estimators", *Empirical Economics*, 29, 115-129.
- Driver, C., and G. Urga, 2004, "Transforming qualitative survey data: performance comparisons for the UK, *Oxford Bulletin of Economics and Statistics*, 66, 71-89.
- European Commission, 1997, "The joint harmonised EU programme of business and consumer surveys", *European Economy* No. 6.
- Folta, T.B., and J.P. O'Brien, 2004, "Entry in the presence of duelling options", *Strategic Management Journal*, 25, 121-138
- Fuchs, V.R., A.B. Krueger, and J.M. Poterba, 1998, "Economists' view about parameters, values, and policies: survey results in labor and public economics", *Journal of Economic Literature*, 36, 1387-1425.

- Guiso, L., and G. Parigi, 1999, "Investment and demand uncertainty", *Quarterly Journal of Economics*, 115, 185-227.
- Hubbard, R.G., 1998, "Capital market imperfections and investment", *Journal of Economic Literature*, 36, 193-225.
- Junankar, S., 1989, "How do companies respond to the CBI's Industrial Trends Survey?", Economic Situation Report, CBI London.
- Junankar, S., 1995, "The CBI's Industrial Trends Survey: its role in monitoring economic trends", Statistics Users' Council Annual Conference, London, February 8th.
- Lamont, O.A., 2000, "Investment plans and stock returns", *Journal of Finance*, 55, 2719-2745.
- Leahy, J.V., and T.M. Whited, 1996, "The effect of uncertainty on investment: some stylised facts", *The Journal of Money, Credit and Banking*, 28, 64-83.
- Mairesse, J., B.H. Hall, and B. Mulkay, 1999, "Firm level investment in France and the United States: An exploration of what have we learned in twenty years", NBER Working Paper No. 7437.
- Mason, R., and H. Weeds, 2001, "Irreversible investment with strategic interactions", CEPR Discussion Paper 3013.
- McDonald, R., and D. Siegel, 1986, "The value of waiting to invest", *The Quarterly Journal of Economics*, 101, 707-728.
- Moel, A., and P. Tufano, 2002, "When are real options exercised? An empirical study of mine closings", *Review of Financial Studies*, 15, 35-64.
- Oulton, N., and M. O'Mahony, 1994, "Productivity and growth: a study of British industry", N.I.E.S.R. Occasional Paper No. 47.
- Pesaran, M. H., 2006, "Estimation and Inference in large heterogeneous panels with a multifactor error structure", *Econometrica*, 74, 967-1012.
- Pindyck, R.S., 1993, "A note on competitive investment under uncertainty", *American Economic Review*, 83, 273-78.
- Proietti, T., 1996, "Persistence of shocks on seasonal processes", *Journal*

of *Applied Econometrics*, 11, 383-98.

Sarkar, S., 2000, "On the investment–uncertainty relationship in a real options model", *Journal of Economic Dynamics and Control*, 24, 219-25

Sarkar, S., 2003, "The effect of mean reversion on investment under uncertainty", *Journal of Economic Dynamics and Control*, 28, 377-396.

Shaanan, J. 2005, " Investment, irreversibility, and options: An empirical framework", *Review of Financial Economics*, 14, 241-254.

Smit, H., and L.Trigeorgis, 2004, *Strategic Investment: Real Options and Games*, Princeton University Press.

Smith,J., and M. McAleer, 1995, "Alternative procedures for converting qualitative response data to quantitative expectations: an application to Australian manufacturing", *Journal of Applied Econometrics*, 10, 165-185.

Trigeorgis, J. 2003, "Real options and investment under uncertainty: what do we know?" in P. Butzen and C. Fuss (eds) *Firms' investment and finance decisions: theory and empirical methodology*, Edward Elgar.

Weeds, H., 2002, "Strategic delay in a real options model of R&D competition", *Review of Economic Studies*, 7, 29-47.

Zarnowitz, V., and L. A. Lambros, 1987, Consensus and uncertainty in economic prediction. *Journal of Political Economy*, 95, 591-621.

---

**TABLE 1**

***PREDICTIONS AS TO WHICH INDUSTRY CHARACTERISTICS WILL OBTAIN FOR EACH CHANNEL OF  
INFLUENCE FROM UNCERTAINTY TO INVESTMENT***

---

Model	Likely sign of uncertainty on investment	Necessary Condition for the sign	Theoretical Variables	Possible proxy variables
REAL OPTION (deferment)	<0	Irreversible investments with the option to postpone and where information is obtained by waiting	High sunk costs; No FMA; High value of waiting	Index of sunk costs; Inverse index of R&D and/or advertising intensity; Persistence in the uncertainty variable
REAL OPTION (expansion or compound)	>0	Option obtained by investment (compound option); alternatively excess capacity minimises stock-out penalties	FMA and/or High cost of non-supply;	Index of R&D and/or advertising intensity Profitability.

---

**TABLE II Results from First Stage Estimation of the Impact of Uncertainty by Industry**

TABLE 1 Results from First Stage Estimation of the Impact of Uncertainty by Industry															
CBI Table <sup>1</sup>	Industry	Standardised Coefficients on Uncertainty Variables <sup>2</sup>									Overall Equation				
		(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)	(12)	(13)	(14)
		coefficient on unc	p-value	sig <sup>3</sup>	coefficient on unc(-1)	p-value	sig <sup>3</sup>	coefficient on unc(-2)	p-value	sig <sup>3</sup>	F-test on unc coeffic- ients <sup>4</sup>	sig <sup>3</sup>	no of obs.	R <sup>2</sup>	F-statistic <sup>5</sup>
24	ferrous metals	-0.2065	0.039	**	-	-		-	-		-		78	0.73	0.000
25	non-ferrous metals	-	-		-	-		-	-		-		78	0.59	0.000
26	building materials	-	-		-	-		-0.1364	0.095	*	-		78	0.79	0.000
27	glass and ceramics	-	-		-	-		-	-		-		78	0.84	0.000
28	industrial chemicals	-	-		-	-		-	-		-		78	0.61	0.000
30	pharmaceuticals and consumer chemicals	-	-		-	-		-	-		-		78	0.52	0.000
32	foundries; and forging, pressing, stamping	-	-		-	-		-	-		-		78	0.76	0.000
33	metal goods nes	-	-		-0.1829	0.034	**	-0.1495	0.095	*	0.0127	**	78	0.83	0.000
35	constructional steelwork	-	-		-0.2858	0.002	***	-	-		-		78	0.73	0.000
37	agricultural machinery	-	-		-	-		0.2483	0.012	**	-		78	0.62	0.000
38	metal working machine tools	-	-		-	-		-	-		-		78	0.67	0.000
39	engineers small tools	-	-		-	-		-	-		-		78	0.72	0.000
40	industrial machinery	-	-		-	-		-	-		-		78	0.62	0.000
41	contractors' plant	-	-		-	-		-	-		-		78	0.73	0.000
42	industrial engines, pumps, compressors	-	-		-	-		-	-		-		78	0.67	0.000
43	heating, ventilating and refrigerating equipment	-	-		-	-		-	-		-		78	0.62	0.000
44	other mechanical engineering	-	-		-	-		-	-		-		78	0.78	0.000
46	electrical industrial goods	-	-		-0.2253	0.058	*	-	-		-		78	0.63	0.000
47	elctronic industrial goods	-	-		-	-		-	-		-		78	0.51	0.000
48	electrical consumer goods	-0.1904	0.036	**	-	-		-	-		-		67	0.75	0.000
49	electronic consumer goods	-	-		0.2074	0.096	*	0.2548	0.035	**	0.0155	**	66	0.60	0.002
50	motor vehicles	-	-		-	-		-	-		-		78	0.73	0.000
52	aerospace and other vehicles	-	-		-	-		0.1937	0.050	*	-		78	0.56	0.000
53	instrument engineering	-	-		-	-		-	-		-		78	0.57	0.000
56	wool textiles	-	-		-	-		0.1714	0.093	*	-		78	0.69	0.000
57	spinning and weaving	-	-		-	-		-	-		-		78	0.67	0.000
58	hosiery and knitwear	-	-		-	-		-	-		-		78	0.50	0.001
59	textile consumer goods	-	-		-	-		-	-		-		78	0.46	0.002
61	footwear	-	-		-	-		-	-		-		78	0.67	0.000
62	leather and leather goods	-	-		-	-		-	-		-		69	0.76	0.000
63	clothing and fur	-	-		-0.2486	0.024	**	-	-		-		78	0.69	0.000
64	timber and wooden products other than furniture	-	-		-	-		-	-		-		78	0.75	0.000
65	furniture, upholstery, bedding	-	-		-	-		-	-		-		78	0.75	0.000
66	pulp,paper, and board	-	-		-	-		-	-		-		78	0.62	0.000
67	paper and board products	-	-		-	-		-	-		-		78	0.61	0.000
68	printing and publishing	-	-		-	-		-	-		-		78	0.54	0.000
70	plastics products	-	-		-	-		-	-		-		78	0.70	0.000

- Notes:
- 1 Table number in CBI Industrial Trends Survey
  - 2 The standardised coefficients multiply the raw coefficient of the relevant independent variable by its standard deviation and then divide this by the standard deviation of the dependent variable
  - 3 \* =significant at 10%; \*\* = significant at 5%; \*\*\* = significant at 1%
  4. F-test of joint significance of sum of uncertainty coefficients where appropriate

**TABLE III Results from Second Stage Analysis**

Estimation by Ordered Probit  
Robust standard errors

Dependent Variable		OPROB1 (1)	OPROB1 (2)	OPROB1 (3)	OPROB1 (4)	OPROB2 (5)	OSUM (6)
Explanatory Variable	Variable Descriptor	sig	sig	sig	sig	sig	sig
Irreversibility	<i>irr</i>	<b>-0.012</b> (0.700)	<b>-0.026</b> * (0.086)				
Value of waiting	<i>persis_opt</i>		<b>-11.233</b> * (0.088)				
	<i>irraug</i>			<b>-0.253</b> ** (0.033)	<b>-0.308</b> ** (0.037)	<b>-0.317</b> ** (0.024)	<b>-0.303</b> ** (0.027)
Opportunities for expansion	<i>rdad</i>				<b>0.543</b> * (0.099)	<b>0.620</b> * (0.067)	<b>0.681</b> ** (0.029)
no. of observations		37	37	37	37	37	37
Prob>Chi <sup>2</sup>		0.483	0.042	0.033 **	0.069 *	0.023 **	0.028 **
Pseudo R <sup>2</sup>		0.01	0.08	0.05	0.11	0.11	0.11

p-values in parentheses

\* = sig at 10%; \*\* = sig at 5%; \*\*\* = sig at 1%



## DATA AND RESULTS APPENDIX

### **A1. Data from the CBI Industrial Trends Survey**

In this paper, we draw upon the Industrial Trends Survey of manufacturing industry carried out by the main employers' organisation, the Confederation of British Industry (CBI). It has been published on a regular basis since 1958 and has been widely used by economists. The survey sample is chosen to be representative and is not confined to CBI members. Questionnaires are targeted at chief executives, managing directors and finance directors. The survey is not confined to CBI members. A core of up to 1100 companies comprise the main panel with up to 300 other new or floating participants. A 50% response rate is typical. (Junankar 1995). The sample is based at the enterprise level except for some of the largest plants where replies are collected for that unit. Our data set is restricted to the period 1978 Q1 to 1999 Q1, since the question on authorisation of investment was added in 1978. The responses in the survey are weighted by net output with the weights being regularly updated and are reported at various aggregate levels. We use the data that are reported for over 40 manufacturing industries, of which we were able to use 37 in our analysis.

The various variables used in the analysis are constructed from the CBI Industrial Trends Survey Questions as follows.

#### *Auth*

Authorisation is interpreted by the vast majority of respondents as board approval. The authorisation data also includes leased assets (CBI 1988, p.29). It is calculated from the balance of replies (% responding more minus % responding less) from Question 3b of the Survey which asks:

Do you expect to authorise more or less capital expenditure in the next twelve months than you authorised in the past twelve months on: plant and machinery? (Possible choices: 'More', 'Same' or 'Less')

#### *opt, unc*

These variables are based upon Question 1 of the survey which asks :

“Are you more, or less, optimistic than you were four months ago about the general business situation in your industry?” (Possible choices: 'More', 'Same' or 'Less')

*opt* is the balance (% responding more minus % responding less) ; for details of *unc* see text.

#### *y, y<sup>e</sup>*

These variables are derived from Question 8 which asks:

“Excluding seasonal variations, what has been the trend over the PAST FOUR MONTHS, and what are the expected trends for the NEXT FOUR MONTHS, with regard to: Volume of output?” (Possible choices: ‘Up’, ‘Same’ or ‘Down’)

The balance of Question 8 approximates to a rate of change in output. Using the generic symbol  $Y$  to refer both to  $y$ ,  $y^e$ , this rate of change is  $\Delta \log Y_t$ , whereas we require  $\Delta \log \Delta Y_t$  as the dynamic output term on the assumption that the log of authorisation is cointegrated with the log of differenced output. It is easy to derive the latter as a Taylor approximation yields:

$$\Delta \log \Delta Y_t = [\Delta \log Y_t + \Delta \Delta \log Y_t]$$

Substituting  $y$ ,  $y^e$  for the LHS of the above equation gives the required definitions as the sum of the relevant survey balance plus the first difference of that balance.

*cu*

This variable is the logit transformation of the % responding ‘No’ to Question 4 which asks

Is your present level of output below capacity (i.e., are you working below a satisfactory full rate of operation)? (Possible choice: ‘Yes’, or ‘No’).

*dcu*

This is the first difference of the variable *cu*.

*fi*

This is the sum of the percentages of respondents reporting either a shortage of internal finance or an inability to raise external finance to *Question 16(c)* of the Survey, which invites respondents to consider which factors, including uncertainty about demand, are “expected to limit capital expenditure authorisations over the next twelve months”. Available replies are:

- inadequate net return on proposed investment;
- a shortage of internal finance;
- an inability to raise external finance;
- the cost of finance;
- uncertainty about demand;
- shortage of labour including managerial and technical staff;
- other.

**TABLE A2 Summary Statistics for Variables in First Stage Estimation**

Variable	CBI Table	Industry	Auth					y					y <sup>e</sup>					opt				
			obs	mean	sd	min	max	obs	mean	sd	min	max	obs	mean	sd	min	max	obs	mean	sd	min	max
24		ferrous metals	86	-10.7	49.4	-92	76	85	-0.9	75.9	-160	220	85	-0.9	66.3	-171	187	86	-23.3	44.4	-95	78
25		non-ferrous metals	86	-1.6	34.0	-72	55	85	-3.5	63.7	-160	140	85	4.3	42.4	-136	90	86	-13.4	30.5	-87	42
26		building materials	86	1.3	28.0	-69	55	85	-3.5	49.3	-112	97	85	3.4	47.7	-103	107	86	-7.5	35.8	-88	57
27		glass and ceramics	86	4.5	31.8	-60	84	85	0.3	40.2	-83	103	85	2.8	37.1	-100	80	86	-7.4	30.3	-74	49
28		industrial chemicals	86	3.0	30.3	-63	84	85	2.9	47.3	-128	108	85	5.1	41.7	-106	92	86	-6.2	33.3	-89	53
30		pharmaceuticals and consumer chemicals	86	5.4	23.9	-51	59	85	13.9	40.3	-109	101	85	20.5	37.4	-72	164	86	-0.2	20.1	-52	37
32		foundries; and forging, pressing, stamping	86	1.6	30.2	-71	67	85	-3.7	42.5	-105	128	85	2.7	36.4	-101	83	86	-7.7	29.0	-78	45
33		metal goods nes	86	1.0	27.5	-71	53	85	-3.6	39.6	-101	74	85	3.9	33.2	-75	78	86	-8.7	27.8	-78	42
35		constructional steelwork	86	-8.4	23.5	-56	46	85	3.5	41.7	-103	105	85	7.3	36.7	-78	118	86	-2.0	29.6	-83	51
37		agricultural machinery	86	-6.8	42.5	-94	84	85	-8.2	64.5	-141	158	85	-4.5	58.7	-186	114	86	-14.8	35.7	-100	59
38		metal working machine tools	86	-5.2	27.9	-65	74	85	-2.1	46.3	-108	110	85	6.1	37.8	-84	104	86	-6.6	31.6	-75	60
39		engineers small tools	86	0.5	29.0	-64	55	85	4.8	47.8	-123	112	85	11.1	37.9	-82	95	86	-2.4	36.2	-78	80
40		industrial machinery	86	-1.7	29.3	-69	58	85	-6.4	53.0	-121	172	85	5.1	47.2	-141	114	86	-7.5	29.4	-90	62
41		contractors' plant	86	-3.8	31.9	-78	69	85	8.1	47.6	-100	152	85	16.2	39.7	-81	127	86	0.2	30.2	-62	68
42		industrial engines, pumps, compressors	86	2.2	28.2	-70	72	85	1.9	41.9	-113	137	85	9.8	32.8	-74	87	86	-5.4	30.6	-88	49
43		heating, ventilating and refrigerating	86	13.8	23.7	-70	62	85	0.2	43.6	-125	82	85	11.0	45.2	-127	122	86	1.3	28.1	-78	48
44		other mechanical engineering	86	-2.4	25.7	-65	47	85	-1.4	39.1	-98	105	85	5.5	33.5	-84	105	86	-3.4	26.5	-68	51
46		electrical industrial goods	86	6.8	31.5	-66	65	85	4.5	51.4	-107	130	85	11.3	58.4	-152	173	86	0.4	33.2	-76	69
47		elctronic industrial goods	86	7.3	31.7	-70	79	85	-0.5	51.3	-136	133	85	9.9	51.0	-143	133	86	-5.4	28.3	-71	54
48		electrical consumer goods	78	6.4	44.3	-88	88	75	-4.8	96.2	-230	219	75	-0.5	76.6	-177	175	78	-17.8	39.7	-91	89
49		electronic consumer goods	82	8.7	36.4	-78	100	77	3.8	73.0	-208	203	77	17.5	55.0	-124	156	82	-3.6	31.9	-69	69
50		motor vehicles	86	0.9	32.9	-64	73	85	-1.7	56.4	-121	158	85	9.7	48.2	-143	114	86	-9.3	33.4	-99	58
52		aerospace and other vehicles	86	2.2	51.2	-86	94	85	6.5	76.6	-228	186	85	6.0	81.7	-258	193	86	-4.5	34.0	-83	87
53		instrument engineering	86	9.4	27.2	-58	67	85	6.6	56.1	-134	130	85	15.8	48.2	-98	140	86	0.0	29.2	-66	56
56		wool textiles	86	-23.1	27.1	-77	49	85	-12.5	44.6	-121	79	85	-9.5	48.4	-123	106	86	-13.4	33.8	-93	75
57		spinning and weaving	86	-5.7	30.7	-81	62	85	-5.7	55.7	-163	135	85	1.4	39.2	-135	85	86	-6.1	36.5	-90	75
58		hosiery and knitwear	86	-2.0	22.0	-55	42	85	1.8	42.8	-115	116	85	7.9	38.0	-113	97	86	-7.5	28.0	-69	44
59		textile consumer goods	86	-8.1	34.3	-72	67	85	4.8	58.7	-145	139	85	6.5	58.9	-174	173	86	-10.3	39.8	-91	67
61		footwear	86	-3.3	29.2	-67	63	85	-7.2	44.9	-108	96	85	0.8	42.0	-131	107	86	-11.6	29.0	-90	55
62		leather and leather goods	80	-8.8	39.9	-83	80	77	-9.2	57.8	-158	108	77	5.4	49.6	-129	100	80	-2.1	33.7	-88	74
63		clothing and fur	86	-7.9	23.8	-60	36	85	0.4	34.0	-72	82	85	9.9	34.9	-92	127	86	-7.2	28.8	-75	41
64		timber and wooden products other than	86	-7.7	28.7	-84	49	85	-1.2	46.8	-108	115	85	5.4	39.9	-89	130	86	-2.0	34.5	-81	81
65		furniture, upholstery, bedding	86	-0.2	28.9	-63	67	85	3.4	44.6	-96	153	85	9.1	40.6	-107	122	86	-4.0	31.5	-89	62
66		pulp,paper, and board	86	-6.9	33.8	-84	72	85	2.7	58.7	-129	131	85	3.6	36.5	-90	82	86	-6.7	37.4	-89	79
67		paper and board products	86	-0.4	31.3	-77	50	85	-1.3	51.9	-129	136	85	9.8	44.8	-90	108	86	-8.0	33.0	-89	73
68		printing and publishing	86	-6.8	22.7	-56	51	85	2.6	31.1	-81	73	85	5.7	36.5	-121	79	86	0.2	27.1	-66	50
70		plastics products	86	10.4	27.1	-59	64	85	6.6	43.7	-107	125	85	13.2	42.4	-80	137	86	-2.1	31.4	-90	65

TABLE A2 continued

Variable		unc					fi					dcu					cu				
CBI Table	Industry	obs	mean	sd	min	max	obs	mean	sd	min	max	obs	mean	sd	min	max	obs	mean	sd	min	max
24	ferrous metals	86	0.30	0.08	0.08	0.465	79	39.5	27.0	2	80	85	0.00	0.43	-0.99	1.176	85	1.40	0.80	-0.70	2.75
25	non-ferrous metals	86	0.34	0.09	0.12	0.472	79	37.9	17.0	0	69	85	0.00	0.47	-1.60	1.806	85	1.27	0.81	-2.20	3.01
26	building materials	86	0.33	0.07	0.17	0.476	79	24.5	13.1	3	67	85	0.00	0.25	-0.75	0.813	85	1.42	0.47	-0.05	2.63
27	glass and ceramics	86	0.37	0.06	0.22	0.469	79	24.1	14.4	2	56	85	0.00	0.28	-0.95	1.380	85	1.48	0.53	-0.95	2.76
28	industrial chemicals	86	0.34	0.07	0.15	0.477	79	28.9	17.4	3	67	85	0.00	0.24	-0.78	0.835	85	1.61	0.42	0.00	2.45
30	pharmaceuticals and consumer chemicals	86	0.31	0.07	0.11	0.449	79	22.4	16.0	0	58	85	0.00	0.15	-0.58	0.567	85	1.71	0.29	0.69	2.34
32	foundries; and forging, pressing, stamping	86	0.38	0.05	0.24	0.477	79	22.3	14.7	3	78	85	0.00	0.19	-0.57	0.535	85	1.46	0.39	0.26	2.08
33	metal goods nes	86	0.39	0.05	0.24	0.476	79	21.2	6.9	7	41	85	0.00	0.16	-0.35	0.406	85	1.47	0.34	0.44	2.10
35	constructional steelwork	86	0.38	0.07	0.20	0.474	79	17.6	9.5	0	49	85	0.00	0.16	-0.43	0.447	85	1.65	0.29	0.87	2.20
37	agricultural machinery	86	0.30	0.09	0.01	0.460	79	32.2	30.2	0	123	85	-0.01	0.63	-2.44	2.447	85	1.36	1.04	-3.04	4.29
38	metal working machine tools	86	0.37	0.06	0.19	0.477	79	21.8	12.7	0	72	85	0.00	0.30	-1.04	1.197	85	1.49	0.54	-1.04	3.00
39	engineers small tools	86	0.37	0.05	0.22	0.471	79	29.3	15.3	6	72	85	-0.01	0.35	-1.78	1.833	85	1.53	0.61	-2.38	3.06
40	industrial machinery	86	0.35	0.07	0.15	0.473	79	22.1	17.1	2	92	85	0.00	0.21	-0.63	0.584	85	1.54	0.38	0.30	2.33
41	contractors' plant	86	0.38	0.07	0.19	0.475	79	13.5	9.6	0	41	85	-0.01	0.24	-0.56	0.497	85	1.48	0.42	0.36	2.23
42	industrial engines, pumps, compressors	86	0.38	0.06	0.15	0.474	79	18.8	11.2	3	67	85	-0.01	0.23	-0.80	0.426	85	1.46	0.46	-0.32	2.11
43	heating, ventilating and refrigerating	86	0.37	0.06	0.21	0.477	79	15.5	10.6	0	42	85	0.00	0.23	-0.66	0.730	85	1.55	0.41	0.04	2.36
44	other mechanical engineering	86	0.38	0.05	0.28	0.456	79	17.7	7.1	5	36	85	0.00	0.14	-0.54	0.512	85	1.50	0.34	0.06	2.00
46	electrical industrial goods	86	0.33	0.06	0.21	0.471	79	21.0	16.6	3	63	85	0.00	0.28	-0.68	0.681	85	1.49	0.50	0.09	2.41
47	electronic industrial goods	86	0.32	0.08	0.14	0.472	79	29.0	22.4	2	80	85	0.00	0.28	-0.89	0.735	85	1.52	0.54	-0.29	2.62
48	electrical consumer goods	78	0.26	0.11	0.01	0.470	79	21.6	20.7	0	83	75	0.01	0.73	-1.75	2.134	75	1.28	1.21	-2.35	3.67
49	electronic consumer goods	82	0.32	0.11	0.01	0.461	79	23.8	22.2	0	89	77	-0.02	0.34	-0.93	0.785	77	1.56	0.55	-0.29	2.61
50	motor vehicles	86	0.32	0.07	0.03	0.473	79	30.7	13.7	1	69	85	0.00	0.26	-0.81	0.802	85	1.38	0.50	-0.51	2.35
52	aerospace and other vehicles	86	0.27	0.11	0.05	0.470	79	35.7	31.0	0	157	85	-0.01	0.43	-1.49	1.176	85	1.45	0.79	-1.49	3.09
53	instrument engineering	86	0.36	0.06	0.16	0.473	79	22.4	16.4	3	75	85	0.00	0.23	-0.61	0.501	85	1.57	0.38	0.43	2.31
56	wool textiles	86	0.37	0.07	0.12	0.472	79	18.3	12.4	0	56	85	-0.01	0.19	-0.65	0.558	85	1.56	0.38	0.46	2.23
57	spinning and weaving	86	0.35	0.07	0.12	0.476	79	19.8	12.3	2	67	85	0.00	0.21	-0.52	0.837	85	1.54	0.41	0.33	2.58
58	hosiery and knitwear	86	0.37	0.06	0.22	0.455	79	15.9	11.7	0	59	85	-0.01	0.15	-0.79	0.314	85	1.72	0.27	0.17	2.13
59	textile consumer goods	86	0.34	0.08	0.09	0.468	79	23.5	17.9	0	75	85	0.00	0.32	-0.82	0.926	85	1.56	0.53	-0.12	2.75
61	footwear	86	0.33	0.08	0.14	0.476	79	19.1	13.9	0	61	85	0.00	0.20	-0.68	0.762	85	1.66	0.34	0.28	2.48
62	leather and leather goods	80	0.37	0.07	0.17	0.476	79	16.7	13.4	0	51	77	-0.03	0.57	-2.37	2.310	77	1.49	0.97	-2.97	4.02
63	clothing and fur	86	0.40	0.05	0.26	0.477	79	22.6	11.0	0	55	85	0.00	0.14	-0.61	0.444	85	1.70	0.25	0.34	2.17
64	timber and wooden products other than	86	0.37	0.07	0.22	0.476	79	18.5	11.0	0	51	85	0.00	0.41	-1.60	1.924	85	1.47	0.71	-2.20	3.25
65	furniture, upholstery, bedding	86	0.37	0.07	0.15	0.471	79	18.8	11.0	0	55	85	0.00	0.37	-2.08	2.158	85	1.58	0.61	-2.68	3.71
66	pulp,paper, and board	86	0.34	0.07	0.12	0.464	79	33.0	18.6	2	81	85	0.00	0.27	-0.78	0.964	85	1.53	0.51	-0.48	2.63
67	paper and board products	86	0.36	0.06	0.16	0.475	79	26.6	12.6	4	60	85	0.00	0.24	-0.78	0.628	85	1.50	0.44	-0.48	2.44
68	printing and publishing	86	0.38	0.05	0.23	0.476	79	22.2	8.6	5	55	85	0.00	0.12	-0.28	0.398	85	1.56	0.26	0.68	2.01
70	plastics products	86	0.38	0.06	0.15	0.475	79	24.8	15.0	6	97	85	0.00	0.17	-0.55	0.541	85	1.59	0.32	0.54	2.36

### **A3. Other Data**

Other data used in the paper are as indicated in the Table

**Table A3**

#### **Additional Data used for Second Stage Analysis**

<b>Variable</b>	<b>Definition and Source</b>
<i>irr</i>	This irreversibility measure was constructed from <i>UK Census of Production</i> data for disposals and acquisitions of plant and machinery (for the period 1979-1989) at the 3-digit level of the 1980 Standard Industrial Classification. The ratio of disposals to acquisitions may be expected to provide a measure of the marketability of second-hand assets. The data for each industry were time averaged over the economic cycle 1979-1989. These 3-digit data were then matched with the CBI industries used for the estimation in this paper. With no strong reason for supposing cardinality, <i>irr</i> was constructed as a reverse ranking of the ratio.
<i>irraug</i>	This used the quartiles of both the <i>irr</i> and the <i>persist_opt</i> distributions (0,1,2,3), summing them, thereby attaching the highest score to industries which were in the highest quartile on both measures (=6). Those in the lowest quartile on both variables had a zero score.
<i>rdad</i>	Indicator variable based upon advertising and R&D intensities; 0=low R&D and low advertising; 1=high on one source but not the other; 2 if high on both. (derived from: Table A2.1 Davies and Lyons 1996)
<i>nprtea</i>	The data used is based on a concordance between the CBI sectors and the 1980 Standard Industrial Classification (1980 SIC) and uses capital stock data kindly supplied by Mary O'Mahony of the National Institute of Economic and Social Research (see Oulton and O'Mahony, 1994). These however were on the basis of the 1968 Standard Industrial Classification (SIC). A correspondence with CBI tables was made using a published reconciliation between the SIC and that for 1980. Profits were calculated from gross value added in each industry less employee compensation and less estimated depreciation in each industry.
<i>OPROB1</i>	Indicator variable for uncertainty significance and sign (-1=negative, +1=positive, 0=null at 10% level).
<i>OPROB2</i>	Indicator variable for uncertainty significance and sign (-2=negative 5% ; -1=negative 10% , +2=positive 5%; +1=positive 10%, 0=null accepted at 10% level).
<i>OSUM</i>	Indicator variable taking on values -2, -1, -1, 0, 1, 2 according to the relative magnitude of the standardized coefficients on the uncertainty variable <i>unc</i> (summed where more than one coefficient is significant) as reported in Table II. All industries with an insignificant coefficient were assigned a zero value. The other 11 industries with significant <i>unc</i> coefficients were given a value of 2 where the sum of the coefficients was greater than an average (across the same set of industries) of the absolute value of the summed <i>unc</i> coefficients. A sign was then attached according to whether the uncertainty effect is positive or negative

**TABLE A4 Summary Statistics for Variables Used in Second Stage Analysis**

CBI TABLE	INDUSTRY	Explanatory variables					Dependent variables		
		irr	persist_opt	irraug	rdad	npratea	OPROB1	OPROB2	OSUM
24	ferrous metals	36	0.10	3	0	-0.03	-1	-2	-1
25	non-ferrous metals	28	0.07	3	0	0.19	0	0	0
26	building materials	30	0.13	5	0	0.30	-1	-1	-1
27	glass and ceramics	31	0.19	6	0	0.16	0	0	0
28	industrial chemicals	37	0.13	5	1	0.10	0	0	0
30	pharmaceuticals and consumer chemicals	33	0.09	3	2	0.40	0	0	0
32	foundries; and forging, pressing, stamping	19	0.16	5	0	0.18	0	0	0
33	metal goods nes	18	0.14	3	0	0.26	-1	-2	-2
35	constructional steelwork	9	0.19	3	0	0.61	-1	-2	-2
37	agricultural machinery	29	0.08	3	2	0.27	1	2	2
38	metal working machine tools	2	0.10	1	1	0.19	0	0	0
39	engineers small tools	3	0.18	3	1	0.16	0	0	0
40	industrial machinery	15	0.09	1	0	0.32	0	0	0
41	contractors' plant	4	0.11	1	0	0.31	0	0	0
42	industrial engines, pumps, compressors	21	0.12	4	1	0.35	0	0	0
43	heating, ventilating and refrigerating equipment	22	0.10	3	1	0.43	0	0	0
44	other mechanical engineering	20	0.17	5	1	0.20	0	0	0
46	electrical industrial goods	25	0.08	2	1	0.20	-1	-1	-1
47	elctronic industrial goods	24	0.11	3	1	0.31	0	0	0
48	electrical consumer goods	32	0.18	6	2	0.16	-1	-2	-1
49	electronic consumer goods	17	0.08	1	2	0.69	1	2	2
50	motor vehicles	35	0.13	5	2	0.18	0	0	0
52	aerospace and other vehicles	27	0.08	2	1	0.20	1	1	1
53	instrument engineering	16	0.12	3	1	0.28	0	0	0
56	wool textiles	7	0.19	3	0	0.07	1	1	1
57	spinning and weaving	5	0.19	3	0	0.03	0	0	0
58	hosiery and knitwear	1	0.15	3	0	0.08	0	0	0
59	textile consumer goods	23	0.10	3	0	0.35	0	0	0
61	footwear	14	0.15	3	0	0.56	0	0	0
62	leather and leather goods	12	0.07	1	0	0.25	0	0	0
63	clothing and fur	10	0.21	4	0	0.58	-1	-2	-2
64	timber and wooden products other than furniture	11	0.12	3	0	0.35	0	0	0
65	furniture, upholstery, bedding	13	0.11	2	0	0.42	0	0	0
66	pulp,paper, and board	34	0.10	4	0	0.06	0	0	0
67	paper and board products	6	0.11	1	0	0.14	0	0	0
68	printing and publishing	8	0.10	1	0	0.39	0	0	0
70	plastics products	26	0.13	4	0	0.30	0	0	0
	mean	19.00	0.13	3.08	0.54	0.27	-0.08	-0.16	-0.11
	standard deviation	10.68	0.04	1.40	0.72	0.16	0.54	0.92	0.83
	minmum	1.00	0.07	1.00	0.00	-0.03	-1.00	-2.00	-2.00
	maximum	37.00	0.21	6.00	2.00	0.69	1.00	2.00	2.00